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# Corrections to “Model predictive control for stochastic max-plus linear systems with chance constraints”, [IEEE Trans. on Aut. Control, 64(1): 337–342, 2019]

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**Abstract**—This paper discusses two issues in connection with [1], namely an error in Proposition 7 and the assumption that the covariance matrix in the chance constraint is positive definite. First we will discuss and correct the error in Proposition 7. Subsequently, we will consider a relaxation of the assumption in Proposition 7 and give a less restrictive and less conservative reformulation of the proposition.

## I. INTRODUCTION

There are two issues in Proposition 7 of [1]:

- In [1] we claim in (17) the following:

$$\lambda_{\min}(\Sigma_z^{-1})\|z - \mu_z\|_2 \leq (z - \mu_z)^T \Sigma_z^{-1} (z - \mu_z)$$

This is based on the erroneous equality  $x^T x = \|x\|_2$ . The correct equality is given by  $x^T x = \|x\|_2^2$ .

- In the proposition we assume that  $\Sigma_z = \Lambda \Sigma_w \Lambda^T$  is positive definite. Note that there holds:  $\text{rank}(\Lambda \Sigma_w \Lambda^T) \leq \min(\text{rank}(\Lambda), \text{rank}(\Sigma_w))$ . However, often the matrix  $\Lambda$  will be a tall matrix, which means that  $\Sigma_z$  will not have full rank (so  $\Sigma_z$  is positive *semi-definite* instead of positive definite), and then for the smallest eigenvalue of  $\Sigma_z$  we have  $\lambda_{\min}(\Sigma_z) = 0$ , which means that the inverse of  $\Sigma_z$  is not defined.

This note is organized as follows. In Section II we give a correction on Proposition 7 in the case where  $\Sigma_z$  is positive definite. In Section III we present an adapted Proposition 7-bis, which is an extension of Proposition 7 and handles the case where  $\Sigma_z$  is positive semi-definite. We also show that the newly derived bound is tighter than the bound derived in Proposition 7 of [1] in the case where  $\Sigma_z$  is positive definite.

## II. CORRECTION OF PROPOSITION 7 OF [1]

In the formulation of Proposition 7, the equation

$$\frac{m}{-\bar{\mu}_z \lambda_{\min}(\Sigma_z^{-1})} \leq \epsilon$$

should be replaced by

$$\frac{m}{\bar{\mu}_z^2 \lambda_{\min}(\Sigma_z^{-1})} \leq \epsilon$$

so the corrected version of Proposition 7 reads as follows:

*Proposition 7:* If  $\Sigma_z$  is a positive definite matrix, let  $\lambda_{\min}(\Sigma_z^{-1}) > 0$  be the smallest eigenvalue of the matrix  $\Sigma_z^{-1}$ . Let  $\bar{\mu}_z(k) = \max_{i=1, \dots, m} \mu_{z,i}(k)$ . If  $\bar{\mu}(k) < 0$  and

$$\frac{m}{\bar{\mu}_z^2 \lambda_{\min}(\Sigma_z^{-1})} \leq \epsilon$$

then

$$\Pr\{\max_{i=1, \dots, m} (z_i(k)) \leq 0\} \geq 1 - \epsilon$$

In the proof of Proposition 7 of [1], equation (17) must be replaced by

$$\lambda_{\min}(\Sigma^{-1})\|z - \mu_z\|_2^2 \leq (z - \mu_z)^T \Sigma_z^{-1} (z - \mu_z)$$

while equation (18) must be replaced by

$$\begin{aligned} \Pr\{\max(z_1, \dots, z_m) \leq 0\} \\ &\geq \Pr\{\|z - \mu_z\|_2 \leq -\bar{\mu}_z\} \\ &\geq \Pr\{\|z - \mu_z\|_2^2 \leq \bar{\mu}_z^2\} \\ &\geq \Pr\{\lambda_{\min}(\Sigma_z^{-1})\|z - \mu_z\|_2^2 \leq \lambda_{\min}(\Sigma_z^{-1})\bar{\mu}_z^2\} \\ &\geq \Pr\{(z - \mu_z)^T \Sigma_z^{-1} (z - \mu_z) \leq \lambda_{\min}(\Sigma_z^{-1})\bar{\mu}_z^2\} \end{aligned}$$

Moreover, equation (19) must be replaced by

$$\begin{aligned} \Pr\{(z - \mu_z)^T \Sigma_z^{-1} (z - \mu_z) \leq \lambda_{\min}(\Sigma_z^{-1})\bar{\mu}_z^2\} \\ \geq 1 - \frac{m}{\bar{\mu}_z^2 \lambda_{\min}(\Sigma_z^{-1})} \end{aligned}$$

The last sentence of the proof of Proposition 7 should be:

If  $\frac{m}{\bar{\mu}_z^2 \lambda_{\min}(\Sigma_z^{-1})} \leq \epsilon$ , therefore, from (18) and (19), we have  $\Pr\{\max(z_1, \dots, z_m) \leq 0\} \geq 1 - \epsilon$ .

As a consequence, the first equation after the proof of Proposition 7 of [1] should be:

$$\mu_z(k) \leq \sqrt{\frac{m}{\epsilon \lambda_{\min}(\Sigma_z^{-1})}}$$

and so equation (20) becomes:

$$\Gamma \tilde{u}(k) \leq \Lambda \mu_w - \Xi(k) + \sqrt{\frac{m}{\epsilon \lambda_{\min}(\Sigma_z^{-1})}}$$

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Finally, equation (21) must be replaced by:

$$\begin{aligned}\Gamma_0(k) &\leq -\Xi_0(k) \\ \Gamma_1(k) &\leq -\Lambda_1\mu_w - \Xi_1(k) + \sqrt{\frac{ms}{\epsilon\lambda_{\min}(\Sigma_{z,1}^{-1})}} \\ &\vdots \\ \Gamma_s(k) &\leq -\Lambda_s\mu_w - \Xi_s(k) + \sqrt{\frac{ms}{\epsilon\lambda_{\min}(\Sigma_{z,s}^{-1})}}\end{aligned}$$

where  $\Sigma_{z,l} = \Lambda_l \Sigma_w \Lambda_l^T$ ,  $l = 1, \dots, s$ .

### III. EXTENSION OF PROPOSITION 7 [1]

In this section we will use the same notation as in [1] and the previous section.

Note that the requirement that  $\Sigma_z$  is positive definite is necessary to be able to use Chebyshev's inequality in the proof of *Proposition 7* in [1]. However, by using the univariate Markov inequality instead of the multivariate Chebyshev inequality we can drop this requirement. The univariate Markov inequality is given as follows:

**Proposition ([2], Proposition 2.6) :** *Let  $Y$  be a non-negative stochastic variable with finite mean  $\mathbb{E}\{Y\}$  and consider a scalar  $\alpha > 0$ , then  $\Pr\{Y \geq \alpha\} \leq \frac{\mathbb{E}\{Y\}}{\alpha}$ .*

The following proposition is an extension of *Proposition 7* in [1] and handles the case where  $\Sigma_z$  is positive semi-definite:

#### Proposition 7-bis:

*Let  $\Sigma_z$  be positive semi-definite. Let  $\bar{\mu}_z(k) = \max_{i=1,\dots,m} \mu_{z,i}(k)$ . If  $\bar{\mu}_z(k) < 0$  and*

$$\frac{\text{tr}(\Sigma_z)}{\bar{\mu}_z^2} \leq \epsilon$$

*where  $\text{tr}(\Sigma_z) = \sum_{i=1}^m [\Sigma_z]_{i,i}$  is the trace of  $\Sigma_z$ , then*

$$\Pr\left\{\max(z_1(k), \dots, z_m(k)) \leq 0\right\} \geq 1 - \epsilon$$

*Proof:* For the sake of simplicity, in this proof, we will write  $z$ ,  $\mu_z$  instead of  $z(k)$ ,  $\mu_z(k)$ . Note that  $\Sigma_z \in \mathbb{R}^{m \times m}$  is symmetric and positive semi-definite (see [1]), so we can write

$$\Sigma_z = U S U^T$$

where  $S \in \mathbb{R}^{\kappa \times \kappa}$  is a diagonal matrix with all positive eigenvalues ( $s_{ii} = \lambda_i > 0$ ) and  $U \in \mathbb{R}^{n \times \kappa}$  contains all eigenvectors corresponding to the positive eigenvalues. Note that due to the fact that  $\Sigma_z = \Lambda \Sigma_w \Lambda^T$  is symmetric,  $U$  can be selected such that all its columns are orthonormal, so  $U^T U = I$ .

Define the signal  $z' = \mu_z + U S^{1/2} v$  where  $v \in \mathbb{R}^\kappa$  is a stochastic variable with  $\mathbb{E}\{v\} = 0$  and  $\mathbb{E}\{vv^T\} = I$ . We

observe that  $\mathbb{E}\{z'\} = \mu_z + U S^{1/2} \mathbb{E}\{v\} = \mu_z + 0 = \mu_z$  and that

$$\begin{aligned}\mathbb{E}\{(z' - \mu_z)(z' - \mu_z)^T\} &= \mathbb{E}\{(U S^{1/2} v)(U S^{1/2} v)^T\} \\ &= \mathbb{E}\{U S^{1/2} v v^T S^{1/2} U^T\} \\ &= U S^{1/2} \mathbb{E}\{v v^T\} S^{1/2} U^T \\ &= U S^{1/2} I S^{1/2} U^T \\ &= \Sigma_z\end{aligned}$$

Therefore, for any  $i, j \in \{1, \dots, m\}$  we have  $\mathbb{E}\{(z_i - \mu_{z,i})^T (z_j - \mu_{z,j})\} = \mathbb{E}\{(z'_i - \mu_{z,i})^T (z'_j - \mu_{z,j})\}$ , and so for any  $i \in \{1, \dots, m\}$  we have  $\mathbb{E}\{(z_i - \mu_{z,i})^2\} = \mathbb{E}\{(z'_i - \mu_{z,i})^2\}$ .

For the expected value of the inner product we now obtain:

$$\begin{aligned}\mathbb{E}\{(z - \mu_z)^T (z - \mu_z)\} &= \sum_{i=1}^m \left( \mathbb{E}\{(z_i - \mu_{z,i})^2\} \right) \\ &= \sum_{i=1}^m \left( \mathbb{E}\{(z'_i - \mu_{z,i})^2\} \right) \\ &= \sum_{i=1}^m \left( \mathbb{E}\{(U_i S^{1/2} v)^2\} \right) \\ &= \sum_{i=1}^m \left( \mathbb{E}\{(U_i S^{1/2} v)(U_i S^{1/2} v)^T\} \right) \\ &= \sum_{i=1}^m \left( \mathbb{E}\{U_i S^{1/2} v v^T S^{1/2} U_i^T\} \right) \\ &= \sum_{i=1}^m \left( U_i S^{1/2} \mathbb{E}\{v v^T\} S^{1/2} U_i^T \right) \\ &= \sum_{i=1}^m \left( U_i S^{1/2} I S^{1/2} U_i^T \right) \\ &= \sum_{i=1}^m \left( U_i S U_i^T \right) \\ &= \sum_{i=1}^m [\Sigma_z]_{i,i} \\ &= \text{tr}(\Sigma_z)\end{aligned}$$

where  $U_i$  is the  $i$ 'th row of  $U$ . Recall that  $\bar{\mu}_z = \max_i \mu_{z,i}$ . Now we derive using Markov's inequality:

$$\begin{aligned}\Pr\{(z - \mu_z)^T (z - \mu_z) > \bar{\mu}_z^2\} &\leq \frac{\mathbb{E}\{(z - \mu_z)^T (z - \mu_z)\}}{\bar{\mu}_z^2} \\ &\leq \frac{\text{tr}(\Sigma_z)}{\bar{\mu}_z^2}\end{aligned}$$

so

$$\Pr\{(z - \mu_z)^T (z - \mu_z) \leq \bar{\mu}_z^2\} \geq 1 - \frac{\text{tr}(\Sigma_z)}{\bar{\mu}_z^2}$$

In Proposition 7 of [1] we derived

$$\begin{aligned} & \Pr\{\max(z_1, \dots, z_m) \leq 0\} \\ & \geq \Pr\{\|z - \mu_z\|_2^2 \leq \bar{\mu}_z^2\} \end{aligned}$$

This means that we have

$$\begin{aligned} & \Pr\{\max(z_1, \dots, z_m) \leq 0\} \\ & \geq \Pr\{\|z - \mu_z\|_2^2 \leq \bar{\mu}_z^2\} \\ & \geq \Pr\{(z - \mu_z)^T(z - \mu_z) \leq \bar{\mu}_z^2\} \\ & \geq 1 - \frac{\text{tr}(\Sigma_z)}{\bar{\mu}_z^2} \end{aligned}$$

□

**Remark:**

Note that in the case that  $\Sigma_z$  is positive definite we find that

$$\begin{aligned} \frac{\text{tr}(\Sigma_z)}{\bar{\mu}_z^2} &= \frac{\sum_{i=1}^m \lambda_i(\Sigma_z)}{\bar{\mu}_z^2} \\ &\leq \frac{m \lambda_{\max}(\Sigma_z)}{\bar{\mu}^2} = \frac{m}{\lambda_{\min}(\Sigma_z^{-1}) \bar{\mu}_z^2} \end{aligned}$$

where  $\lambda_i(\Sigma_z)$  is the  $i$ -th eigenvalue of  $\Sigma_z$ . This means that when  $\Sigma_z$  is positive definite the bound in Proposition 7-bis is in general tighter than the one derived in Proposition 7 of [1].

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